Analysis of AHP by BIBD

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ABSTRACT

The essence of AHP is to evaluate objects in terms of the eigen yector of the comparison-matrix. But when the number of objects, n_s is too large, it causes often worse reliability for an observer to evaluate all paired comparisons at a time. So it is necessary to decompose the whole set of pairs into several classes, and for each class to be evaluated by one observer. We propose the decomposition by BIBD (balanced incomplete block design) well known in the field of experimental design or combinatories. We show by simulation experiments that our method gives better evaluations than the ordinary AHP

In connection with these, we propose the logarithmic least square method, very easy to calculate, and show that this gives very good approximation to the eigen vector method when n is rather small, and that the former completely coincides with the latter when $n \le 3$, surprisingly

§1 Introduction

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The essence of AIIP (Analitic Hierarchy Process) is to evaluate objects in terms of the eigen vector corresponding to the maximal eigen value of a matrix whose (i, i) element is the ratio of evaluation of object j to object j [1] [2] [3]. The idea is to intend to unify local informations taken by paired comparison into a gloval information.

But when the number of objects is too large, it causes often worse reliability for a observer to evaluate all paired comparisons at a time. In such case, it is necessary to decompose the whole set of paires of objects into several blocks, and for each block to be observed by one observer. It is important how to decompose the set of pairs. We propose the decomposition by BIBD (Balanced Incomplete Block Design) well known in the field of experimental design [4] [5]. And we show that this method is to give better evaluation under certain assumptions by simulation experiments (§4). Further we propose the logarithmic least square method for our problem, and show that this gives a good approximation to the eigen value method in AHP (§3).

§2 BIBD

Let $E = \{1, 2, ..., v\}$ be the set of objects i = 1, 2, ..., v. The number of pairs among E is ${}_{i}C_{2} = v(v-1)/2$, and if v is large this becomes very large and an observer cannot compair all such pairs at a time, let the set of objects which an observer can accommodate with sufficient reliability, be called an "allowable block", and let us denote the size of allowable block by k ($\leq v$).

Then we need to decompose the set of ${}_{n}C_{2}$ pairs into the classes of size ${}_{n}C_{2}$ and to allocate several observers to these classes. Each observer makes paired comparisons in his class. We unify these results and can get the evaluation on E_{n}

For example, there are v=7 applicants for a prize essay, and we try to judge their essays and to decide ranking on them. Let the allowable block size be k=3, that is, one judge can read 3 essays and make paired comparisons on them. In this case, the ${}_{7}C_{2}=21$ pairs are decomposed into classes of size ${}_{3}C_{2}=3$. So we need ${}^{2}21/3=7$ judges. We unify the results of 7 judges into the whole ranking on 7 applicants.

It is the problem how to decompose the set of ${}_{\nu}C_{2}$ pairs into blocks and how to unify the results of observations on blocks into the whole evaluation. We propose the decomposition by BIBD and the unification by the eigen value analysis used in AHP

BIBD on $E=\{1,2,\dots,r\}$ is the class $D=\{B_1,r\}$ By $\{1,6\}$ of subsets (called "blocks") $B_i\subseteq E$ (i=1-6) satisfying the followings.

 $(B_t)^{\dagger} = k$ (the block size is constant)

for any t=1-x $\{h_t\}_t\in B_{p_t}t=1, h_t = t \} \text{ (the repetition number is constant)}$

for any $k, j \in I$, $k \neq j$ If $k \in B_k$, $j \in B_k$, k = 1, by k if the intersection number is constant)

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(where | S | denotes the size of a set S.)

It is clear that (iii) implies (ii), so (i) and (iii) suffice for D to be BIBD. Specifically a BIBD with $\lambda = 1$ is called Steiner system and here we consider only Steiner system, which is denoted by (v, k)-D.

Example 1 The class of subsets of $E=\{1,\ 2,\ ...,\ 7\}$ shown in Table 1 is $(7,\ 3)$ -D. The pairs in each block are shown in write hand of the block. All such pairs construct the set of $_7C_2$ pairs. In other word, the set of pairs in E is decomposed into pairs in blocks.

We can represent this situation in terms of graph theory in the following way: To construct (v, k)-D is equivalent to decompose the set of edges of a complete graph with v points into complete graphs with k points (-Fig. 1)

	Table 1 (7, 3)-D							
B. =	$[1, 2, 4] \rightarrow 12, 14, 24$							
B ₂ =	$ 2, 3, 5 \rightarrow 23, 25, 35$							
$\vec{B_3} =$	[3, 4, 6] - 34, 36, 46							
$B_4 =$	$[4, 5, 7] \rightarrow 45, 47, 57$							
$B_5 =$	$ 5, 6, 1 \rightarrow 56, 51, 61$							
ß, =	$ 6, 7, 2 \rightarrow 67, 62, 72$							
<i>II-</i> =	$ 7, 1, 3 \rightarrow 71, 73, 13$							

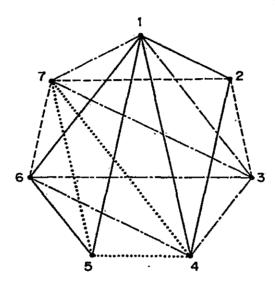


Fig. 1 BIBD decomposition

By the graph theoretic representation staded in Ex. 1 we can easily have the following relations.

$$vr = bk$$
 (2 · 1)
 $b = \frac{vC_2}{kC_2} = \frac{v(v-1)}{k(k-1)}$ (2 · 2)

If (v, k) D exists then integers v and k satisfy (2, 1), (2, 2) with integers v and k, so for any integers v and k we do not necessarily have (v, k).D. For example (8, 3) D never exists, But (9, 3) D exists, so we can treat the case v = 8, k = 3 by taking one of objects in (9, 3) D as dummy.

The conditions of existence and construction methods of (v, k)-D have been widely and deeply researched in the field of experimental designs and combinatorial theories [4] [5].

In order to show why the decomposition by the (v, h)-D is appropriate, we propose another rather natural decomposition shown in Table 2. Of course this $D = |B_1, ..., B_7|$ is not BIBD, where pair (1, 2) occurs 2 times in B_1 and B_7 , while pairs (1, 4), (1, 5) do not occur anywhere.

The author believe that $(v, k)\cdot D$ would give the best possible decompositions for our problems. In the end of this section we give another $(v, k)\cdot D$ in Table 3.

		anie z	
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$B_1 =$	11, 2, 3	 12, 13,	23
$B_2 =$	2. 3. 4	→ 23, 24,	34
$B_3 = .$	13, 4, 51	→ 34, 35,	45
$B_1 =$	[4, 5, 6]	45 , 46.	56
$B_5 =$	15, 6, 71	→ 56, 57,	67
B ₆ =	16, 7, 11	→ 67, 61.	71
B- =	17. I. 2i	→ 71. 7 1	12
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§3 LLS, AllP estimation

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Let $[1, \cdots, n]$ be the set of ofjects $i = 1, 2, \cdots, n$. An observer observes the ratio of evaluation of object i to object j and let us denote the observation by x_{ii} .

We assume the statistical model of xi to be

$$x_{ij} = a_{ij} \cdot e_{ij}, \quad a_{ij} = w_i/w_j \qquad (i < j, i, j = 1 - u)$$

$$x_{ij} = b/x_{ij} \qquad (3 \cdot 1)$$

Where w_i (>0) is the real evaluation of object i and is an unknown parameter, and e_{ij} (>0) is an idependent random variable representing the error of the observation. And we always recognize that any multiples of $\{w_1 \cdots w_n\}$ are equivalent to $\{w_1 \cdots w_n\}$ itself. Further we assume that

$$E(lne_{ii})=0, \qquad V(lne_{ii})=\sigma^2(n) \tag{3.2}$$

and $\sigma^2(n)$ is a monotone increasing function of n, the number of objects to be observed.

Whether these assumptions are reasonable or not is a psychological or a physiological problem, but we can agree with these assumptions as a trial scheme.

The main purpose of AHP is to get estimates \dot{w}_i of w_i (i = 1 - n)by calculating the eigen vector $\dot{\mathbf{w}} = [\dot{w}_1, \cdots, \dot{w}_n]$ corresponding to the maximal eigen value λ of the nxn comparison matrix $\mathbf{X} = [x_{ij}]$.

Of course we have other estimation methods. The most natural one is "logarithmic least square (LLS)". For simplicity let $x_{ij} = \ln x_{ij}$, $w_i = \ln w_i$ and $c_{ij} = \ln c_{ij}$. Then we have

$$\widetilde{x}_{ij} = \overline{w}_i - \overline{w}_j + \widetilde{\epsilon}_{ii} \quad (i < j, i, j = 1 - n) \tag{3.3}$$

Appling the least square method to (3 · 3) we have least last square estimate \vec{w}_i of \vec{w}_i , and taking inverse transform we have $\vec{w}_i = e^{\vec{w}_i} (i = 1 - n)$. This is LLS estimation.

For example let n=3. Then we have

$$\vec{x}_{12} = \vec{w}_1 + \vec{w}_2 + \vec{e}_{12}, \ \vec{x}_{13} = \vec{w}_1 - \vec{w}_2 + \vec{e}_{13}, \ \vec{x}_{23} = \vec{w}_2 - \vec{w}_3 + \vec{e}_{23}$$
 (3 · 4)

As the vector $\mathbf{w} = [w_1, \cdots, w_n]$ multiplied by an arbitral constant is equinalent to \mathbf{w} itself. We can assume $\mathbf{w}_1 \mathbf{w}_2 \mathbf{w}_3 = 1$, so we have

Appling least square method to(3 * 4),(3 * 5)and taking inverse transform we have

$$\hat{\mathbf{w}}_1 = (\mathbf{x}_{12} \ \mathbf{x}_{13})^{1/3}, \quad \hat{\mathbf{w}}_2 = (\mathbf{x}_{23} \ \mathbf{x}_{23})^{1/3}, \quad \hat{\mathbf{w}}_3 = (\mathbf{x}_{32} \ \mathbf{x}_{32})^{1/3}$$
(3 · 6)

This estimation is very simple, but we have a surprising fact that \underline{w}_i (i = 1 - 3) in (3 · 6) coincide with the components of the eigenvector corresponding to the maximal eigenvalue of the comparison matrix $X = [x_n]$. That is, we have

Theorem 1 In the case $n \le 3$, LLS estimates coincide with AHP estimates. Proof In case of n=2, we have easily

$$\dot{w}_1 = \sqrt{x_{12}}, \quad \dot{w}_2 = \sqrt{x_{21}} \tag{3.7}$$

as the LLS estimates of w1, w2 respectively. And by direct calaculation we have

$$X = \begin{bmatrix} \overline{w}_1 \\ \widetilde{w}_2 \end{bmatrix} = 2 \begin{bmatrix} w_1 \\ \widetilde{w}_1 \end{bmatrix} \qquad X = \begin{bmatrix} 1 & x_{12} \\ x_{21} & 1 \end{bmatrix}$$

and the maximal eigen-value of 2×2 comparison matrix is 2, so $\hat{w}_1\hat{w}$, in (7) are the AHP estimates. In case of n=3, from (6) we have

$$\begin{bmatrix} 1 & x_{12} & x_{13} \\ x_{21} & 1 & x_{23} \\ x_{33} & x_{32} & 1 \end{bmatrix} \quad \begin{bmatrix} \hat{w}_1 \\ w_2 \\ \hat{w}_3 \end{bmatrix} \quad = \lambda \quad \begin{bmatrix} \hat{w}_1 \\ \hat{w}_2 \\ \hat{w}_3 \end{bmatrix}$$

$$\lambda = 1 + r + 1/r$$
, $r = \sqrt[3]{x_{12} x_{23} x_{31}}$

by direct calculation. But from Perron & Frobenius Theory we can state that if a positive matrix has a positive eigen vector then it corresponds to the maximal eigen value, (A positive matrix (vetor) means a matrix (vector) whose components are all positive), (See the proof of the theorem 4 in 16th.

Like the case n=3 we have general LLS estimates

$$\hat{w}_i = (\prod_{i=1}^n x_{ii})^{1/n} \qquad i = 1, \dots, n \tag{3.8}$$

which can be stated simply as w, is the geometric mean of i-th row of the comparison matrix X

But unfortunately they no longer coincide with AHP estimates for n>3. Thus if n>3 then Theorem 1 does not hold. But through Theorem 1 we can state that if $\sigma^2(n)$ is reasonably small LLS estimates must be good approximations to AHP estimates even if n is greater than 3. This also teaches us that our statistical model (3 · 1) is valid for the AHP analysis.

§4 Decomposition Methods by BIBD

Now we propose our decomposition methods by BIBD, We are given the set of objects $E=\{1,2,\cdots,n\}$ to be evaluated. Let the allowable block size k be far smaller than v. We decompose E into blocks B_1,\cdots, B_k which construct Sfeiner system (v, k)-D.

Step 1 For each $B_r = [\beta_1, \beta_2, \cdots, \beta_k] \subseteq E$ an observer observes the objects and gets observation x_{β_1,β_2} by a paired comparison β_1 to β_2 (t<s; t, s=1, ..., k). Note that the observation error of x_{β_1,β_2} (measured by V (in x_{β_1,β_2}) = $\sigma^2(k)$) is far smaller than the one incurred by the observation in the whole set E. Step 2. Let

$$S_c = \mathrm{i} x_{\mu_1,\mu} - x_{\mu_1,\mu} - \cdots, x_{\mu_{n-1},\mu}$$

and let

$$\begin{split} S_i &= \left[\mathbf{x}_{[\theta_i,\theta_i]} \mathbf{x}_{[\theta_i,\theta_i]} \cdots \mathbf{x}_{[\theta_i,\theta_i]} \right] \\ &= \mathbf{B}/\mathbf{x}_{[\theta_i,\theta_i]} \mathbf{1}/\mathbf{x}_{[\theta_i,\theta_i]} \cdots \mathbf{1}/\mathbf{x}_{[\theta_i]=\theta_i} \right] \end{split}$$

for e=1 2, ...b

Then by the properties of BIBD

$$(S_1, S_1)^t (S_1^{-t} S_0^{t-t})^t = (S_0^{-t} S_0)$$

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constructs the set of all paired comparisons x_n (r > 1 -r), from which we construct the $v \times v$ comparison matrix $X = [x_n]$

Step 3 We apply the usual AHP to the comparison matrix X, that is, we calculate the eigen vector $\hat{\boldsymbol{w}} = [\hat{\boldsymbol{w}}_1, \hat{\boldsymbol{w}}_2, ..., \hat{\boldsymbol{w}}_r]$

corresponding the maximal eigen-value of X. The \hat{w}_{s} is the desired estimate of \hat{w}_{s} in (3 + 1). Example 3 (v=7 k=3, BIBD decomposition)

Let $\mathbf{w}_1, \dots, \mathbf{w}_7$ (3 · 1) take the values shown in Table 4. Of course these are unknown for the observers and are to be estimated. Further $a_n = \mathbf{w}_i/\mathbf{w}_r$ (i. $j = 1 \sim 7$) are also shown in Table 4.

For each block B_c (shown in Tabli 2) an observer takes observations $x_{d_1d_2}, x_{d_1d_3}, x_{d_2d_3}$ where

$$x_{i_1 + i_2} = a_{i_1 + i_3} e^{i_1 + i_2} \tag{4.3}$$

and $e_{A_2A_3}$ is a random number whose logarism in $e_{B_1B_2}$ normally distributes with zero mean and variance σ (3) by our model (3 ° 1), (3 ° 2) (For the actual value of σ (3) see (1 ° 5)). For $e=1, 2, \cdots, 7$, $S_e=\lim_{t\to a_1B_2} s_{A_1B_2}s_{A_$

$$X_{c} = \begin{bmatrix} 1 & x_{d_{1}d_{2}} & x_{d_{1}d_{2}} & x_{d_{1}d_{3}} \\ x_{d_{2}d_{1}} & 1 & x_{d_{2}d_{3}} & 1 \\ x_{d_{3}d_{1}} & x_{d_{2}d_{2}} & 1 \end{bmatrix}$$
(1 - 1)

Unifying X_1, \dots, X_7 we have the 7×7 comparison matrix X shown in Table 6, and calculating the eigenvector for the maximal eigenvalue λ of X we have the estimates $\mathbf{w}_1, \dots, \mathbf{w}_7$ shown also in Table 6. Comparing \mathbf{w}_i to \mathbf{w}_i (i = 1 - 7) we can say that we have generally fairly good estimates.

Now we consider the actual value of $\sigma^2(n)$ in our model (3 * 1). Of course this depends on the given real problem. But we assum

$$\sigma(3)^2 = 0.158^2$$
, $\sigma^2(4) = 0.250$, $\sigma^2(7) = 0.5^2$, $\sigma^2(13) = 0.980^2$ (4 · 5)

as trial values in our simulations, where $\sigma'(n)$ is roughly proportional to ${}_{n}C_{2}$. Example 4 (v = 7, direct method)

Here we will describe the usual AIIP method applied directly to Table 4. We multply a_{ij} by a random number e_{ij} and have

$$s_{ij} = a_{ij} c_{ij}$$
 $(i < j, i, j = 1, \dots, 7)$

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where $\ln c_n$ normally distributes with zero mean and variance $\sigma^2(7) = 0.5$ (-(4.5)), and calculate $x_n = 1/x_n$, then we have the comparison matrix $X = [x_n]$ shown in Table 7. Calculation the eigen vector for the maximal eigen value λ we have estimates \hat{w}_i ($i = 1, \dots, 7$) shown alson in Table 7, quite worse than the ones in Table 6.

Next we try to investigate another case $\sqrt{-13}$, k=4 in Examples 5, 6 along the same line as above. Example 5 (v=13, k=4, BIBD decomposition)

We show w_i (i = 1 ~ 13) and $a_{ij} = w_i/w_i$ (i, j = 1 ~ 13) in Table 8, and (13, 4) – D and its comparison abservations in Table 9, where the logarism of the random members have the variance $\sigma^2(4) = 0.250^2$ (in (4 · 5)). Fimally the unified comparison matrix X and its maximal eigen value and the corresponding eigen vector are shown in Table 10. These give us rather good estimates.

Table 4 7×7 [a ₁₄]								
		1	2	3	4,	,5	6	7
w _i =1	1	1:	1.857	20.43	8.094	8.004	3.605	4.516
$w_2 = 0.538$	Z	0.538	L	11.00	4.358	4.358	7,941	2.432
w3=0.049	3	0.04895	-0.09091	3	0.3962	0.3962	0.1765.	0.2211
w ₄ ⇔0.124	4	G 1235	0.2294	2.524	1	1,	0.4154	0.5579
w ₅ ≃0.124	5	0,1235	0.2294	2.524	J	1	0.4454	0,5579
w<=0.277	6	0.2774	0,5152	667	2 245	2.245	ī	1.253
my=0.221	7	0 2214	0.4113	1 521	1 792	1.792	0.7983	1

			Table 5	Ar X	V					
8 ₁ = 11,2, X ₁ =	1 41 1 1 2 0.6493 4 0.138		236 269	B ₂ =	$X_2 = 3$	2 1 0.07665 0.1621	3 5 13.05 6,171 1 0.405 42,465 1	e -		
B ₁ = [3,4, X ₅ =	3 7	4 6 0.4033 0.	1975	以 4≃		4 1 1 6 0.9709	S / L.030 0,559 L. 7,551			
B, 15 6, R		1 0 : 3,800 1 1 3	1135 2632 361	I L; =	6.7.2 I	(1 7×9	7 · 1,226 0,531	ī	3 ^	
X, ≈,	$X_{n} = A - V \begin{bmatrix} 5.347 & 1 & 19.33 \\ 3 & 0.2293 & 0.05173 \end{bmatrix} $ $X_{n} = -7 \begin{bmatrix} 0.8160 & 1 & 0.3623 \\ 2 & 1.884 & 2.760 & 1 \end{bmatrix}$ Table 6 $7 \times 7 \times = \{x_{n}\}$ (BBD decomposition) 1 2 3 4 5 6 7 \times $\lambda \div 7.0291$									
3 0, 4 0, 5; 0, 6 0,	1,540 ,6493 1 ,05173 0,0766 ,1382 0,2342 ,1135 0,1621 ,2632 0,5317 ,1870 0,3623	2.479 2.465 5.063	7,236 4,269 0,4033 1 0,9709 3,904 1,789	8,814 6,171 0,4057 1,030 1 2,572 1,804	3,800 1,881 0,1975 0,1129 0,3887 1 0,8160	5,347 2,760 0,2293 0,5590 0,5544 1,226 1	₩ ₂ =0,6067 ₩ ₄ =0,0198		•	
1	2,			[x ₁₁] (usua 5,	i Altr)	7	λ =7.323		ŝ.	
0.5	1.731 777 1 7524 0.08562 259 0.2119 118 0.1601 959 0.4570	13.29 11.68 1 0.707 0.8402 6.262 3.031	7.942 4.720 0.3694 1 1.801 2.688 2.569	8,944 6,245 1,190 0,5543 1 1,714 1,513	1.437 2.188 0.1597 0.3720 0.5835 1	4.561 4.173 0.3299 0.3892 0.6611 2.150	$w_1 = 1$ $w_2 = 0.7679$ $w_3 = 0.0793$ $w_4 = 0.1303$ $w_5 = 0.1460$ $w_6 = 0.4139$ $w_1 = 0.2201$	ŗ	•	
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Example 6 (v=13, direct method)

If an observer observes a_0C_2 paired comparisons x_0 (i < j; i, j, = 1 - 13) at a time, then the logonithm of observation error e_0 has variance σ^2 (13) = 0.980 2 (\rightarrow (4 · 5)). We calculate $x_0 = a_0$ e_0 from Table 8 and the random numbers e_0 with above mentioned properties, $X = |x_0|$ and its meximal eigenvalue and its eigenvector are shown in Table 11. These estimates almost have no reliability.

At the end of this section we note that we can use LLS estimation (-\$3) for our purpose.

First we apply LLS estimation (\rightarrow (3 * 8)) to Table 6 (Example 3) and have \dot{w}_i as estimate of w_i ($i=1\sim7$) (in Table 4). These are shown in Table 12, where w_i (in Table 6) are shown again for the comparison. (Of course \dot{w}_i 's are standardized as $\dot{w}_1'=1$). We can see that \dot{w}_i gives a surprisingly good approximation to w_i ($i=1\sim7$).

Second appling LLS to Table 7 we have Table 13. This also gives farely good approximations. Next we have Table 14 from Table 10 and Table 15 from Table 11 by the same way as above.

Generally we can say that the LLS estimation gives a good approximation to the eigen vector estimation when the abservation error is small and the number of abjects is small.

The labor of the calculation of the LLS is far easier than the eigen vector method. The former is easily done on a desk calculator of pocket size, but the latter needs at least a personal computer So the advantage of LLS method should be highly appreciated even in the highly computerized countries like Japan 11SA and etc. Never the less the author thank that the deep meaning of AHP is concealed in using the eigen vector for a estimation

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Table 8 13 * 13 Gt
                  2 3,1 5,6 7
3.511
                                                                                          3, 187

    w<sub>A</sub>
    =0.1086
    3
    0.81
    0.9
    1
    1.111

    w<sub>A</sub>
    =0.0977
    .4
    0.81
    0.9
    3
    1.111

    w<sub>A</sub>
    =0.0880
    5
    0.6561
    0.729
    0.81
    0.9

w.s =0.1086 3 0.81 0.9
                                 1,524 1,693 1,882 2,091
                                                                                          2.323
                                                                            1.521 1.693 1.882
w. =0.0641 8 0.4783 0.5314 0.5905 0.6561 0.729 0.81 0.9
                                                                 1.111 1.235 1.372 1.524 1.693
                                                          1
1.524
wa=0,6168 11 0,3187 0,3874 0,1365 0.1783 0.3[1 0.3]0 0.6361 0.729 0.81 0.9
                                                                            1
                                                                                   1 111 1 235
wi--0.0021 12 0.3138 0.3187 0.854 0.406 0.483 0.314 0.500 0.666 0.720 0.81 0.00
w_0 = 0.0379 - 13 + 0.2824 - 0.3138 - 0.3187 - 0.3874 - 0.4305 - 0.4783 - 0.514 - 0.5005 - 0.6564 - 0.729 - 0.84 -
                                   Eth v v v adal c
                 t
                        m
                                                                      3 3

        1
        1
        1
        2.72
        1.057
        0.111
        2
        1
        0.9062
        2.113
        0.06

        2
        0.7862
        1
        0.9282
        1.115
        \
        0.105
        1
        1.088
        1.666

        4
        0.9461
        1.077
        d
        2.648
        \
        5.06733
        0.7265
        1
        6.95

                                                                    324 1 167 1,056 2,407 (
                                                                    1 0.869 1 0.8383 2.667 1
6 0,9179 1,193 1 1.951
    10 0.3181 0.7067 013776 1
                                    11 0.3262 0,6002 0,5163 1
                                                                     2 0.4155 0.375 0.5118 1
                                                                       6
    2 1.398 1.744 2.697 1
                                       ×
                                             9
                                                   115
   3 1.867
                                            12 1
                                       11
           11
                                                                       12.
                                                                            413
9 1,177 0,5362 0,5076 1
                                   Table 9 X<sub>1</sub> X<sub>2</sub>
                                                 Art (continued)
      1,0000 1,2720 1,7500 1,0570 0,9139 1 7120 2,2150 1,9590 1,8650 3,1440 4 4770 2,1470 4,9400
       0.7862 1.0000 0.9462 0.9282 2.1130 1.3980 1.7440 1.6910 2.6970 1.4150 3.0660 3.7910 4.0540
       0.5682 1.0570 1.0000 1 1670 1.3880 1.0550 1.8670 1 3240 1.9700 1.3850 1.6660 2.4070 2.9660
       0.9461 1.0770 0.8569 1.0000 1.1890 0.8383 1.5500 1.0120 1.7320 2.6480 2.0870 2.6670 1.5230
       1.0940 0.4733 0.7205 0.8410 1.0000 1.3280 0.7647 1.5230 1.3660 1.6130 1.9370 1.7690 1.9490
       0.5841 0.7153 0.9479 1.1930 0.7530 1.0000 1.1350 1.1470 1.9040 1.1620 2.5040 1.9540 2.9150
       0.5105 0.5914 0.7553 0.9597 0.6566 0.8718 0.9285 1 0000 1.5530 0.9633 0.6957 1.2680 1.6950
       0.5362 0.3708 0.5076 0.5774 0.7321 0.5252 0.8696 0.6439 1.0000 0.9799 1 1470 1.4210 1 1751
       0.3181 0.7067 0.7220 0.3776 0.6200 0.8606 1.0340 1.0380 1.0210 1.0000 0.9447 1.1700 1.4540
       0.2237 0.3262 0.6002 0.4792 0.5163 0.3994 0.5141 1.437c 0.8718 1.0590 1.0000 0.9362 1.1010
       0.4658 0.2638 0.4195 0.3750 0.5653 0.5118 0.6242 0.7885 0.7037 0.8547 1.0680 1.0000 0.9769
       0,2024 0,2467 0,3372 0,6568 0.5132 0.3431 0,5417 0.5901 0.8497 0,6876 0.9082 1.0240 1.0000

    eigen value - 13.343170

      • sector
       0.00000-0.88277-0.71639 \pm 0.70766-0.60639-0.62915-0.49856-0.566, \\ -0.38128-0.18178-0.33566-0.31258-0.27813
```

Ô

ð

Table 11 13×13 $\lambda = \{x_n\}$ (usual AHP)

	:	2	3	4	5	6	7	8	9	10	H	12	13
ı	1	1 774	1.071	2.175	1.919	0.5758	1.045	12.48	16.1 6	3,908	0.8363	1 902	3,884
ļ	0,5636	1	0.8771	3,227	1.781	0.4639	0.2651	1.411	1.370	0.9083	1.924	8,971	3,072
ŀ	0.9338	1 140	1	6.522	1.000	2,094	0.4587	4.016	3.730	1.012	3,680	0.6331	2.824
١	0.4599	0.3099	0.1533	t .	7,611	0.9521	1.959	0,1990	0.2546	0,8926	2.651	0,5812	3, 156
١	0.5212	0.5615	0.999	0.1314	1	1.357	1.393	1.208	0.8352	1.392	8,979	6.493	0 7722 *
1	1.737	2.156	0.4776	1.050	0.7372	1	0.6666	2,710	8.762	10.76	0.6958	1.390	3, 136
1	0.9574	3.772	2,180	0.5104	0.7178	1.500	1	1,085	0,7848	2,004	t4.61	6.686	3.485
1	0.08013	0.7087	0.2490	5.024	0.8277	0.3691	0.921	1	1.584	0.5654	6,7;tt	1.110	7.981
Ì	0.06187	0.7300	0.2681	3.927	1.197	0.1141	1.274	0,6312	ı	1.145	1.339	0.5351	2,335
٠		1.101	0.9879					1.889	0,8730	1	2.956.	0.8129	1 162
	1.196	0.5197	0.2748	11,3772	0,114	1.437	0.06012	0.1186	0.7115	0.3383	1	0.8814	1.807
ŧ													0.05509
ì	0.2574	0.3255	0,3541	0.2894	1.295	0.2911	0.2869	0,1253	0.1283	0,8563	0.5535	18 1.	ı

* rigen value=19.80(88

• version ==

Pals	Calde 12 (Decomposition)		2 Table 13			Table 11			Çable 15		
(Dre			(AHP)		,	(the composition)			ĩ		
ı	ú.	ú,	,	ú,	ē,		ú,	ů,		ü,	н,
- 16	T.	1	t		1	- 1	t	1		1	1
2	0,6067	0.0061	2	0.7679	0.78601	2	10,820236	O.W.Bis	2	01 \$795	04002
3	0.0496	0,0498	3	0.0793	0.0755	3	0.7164	0.7354	ı	0.6115	0144017
4	0.1222	0.1258	4	0.1303	0.1273		0.7076	0.7126	1	0.4551	15,76,280
5	0.1136	0.1137	5	0.1460	0.1385	5	0.6064	0.6000	5	0.4517	0.5139
6	0.2916	0.2904	6	0.4139	0,4067	6	0.6292	0.6439	6	0.6325	0.7705
,	0.2131	0.2131	7	0.2201	0.2224	7	0.4906	9,5192	7	0.7016	0.0128
	(cigen	(1.1.5)		(rigen	(LLS)	×	0.4643	0.1723	11	0,5000	0.1611
	ALCAN)		vector)		9	0.1813	0.3926	4	0.3004	0.3210
						10	0.4148	0.4206	[4r	0.2703	0.3635
						11	0.3356	0,3354	11	0.2202	0.2123
						12	0.3126	0.3269	12	0 2179	0.21%
						13	0.2781	11.2833	13	0.3613	0.2544
							frigen Arrivet	(LIS)		frigen Vertigen	(1.18)

Conclutions

For the statistical model stated in $(3 \cdot 1)$ $(3 \cdot 2)$ our BIBD methog gives better results than the usual AHP method. Further the LLS method $(3 \cdot 8)$ gives very good approximation to the eigen value method.

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